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# Research article

# The extended Weibull–Fréchet distribution: properties, inference, and applications in medicine and engineering

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**Abstract:** In this paper, a flexible version of the Fréchet distribution called the extended Weibull– Fréchet (EWFr) distribution is proposed. Its failure rate has a decreasing shape, an increasing shape, and an upside-down bathtub shape. Its density function can be a symmetric shape, an asymmetric shape, a reversed-J shape and J shape. Some mathematical properties of the EWFr distribution are explored. The EWFr parameters are estimated using several frequentist estimation approaches. The performance of these methods is addressed using detailed simulations. Furthermore, the best approach for estimating the EWFr parameters is determined based on partial and overall ranks. Finally, the performance of the EWFr distribution is studied using two real-life datasets from the medicine and engineering sciences. The EWFr distribution provides a superior fit over other competing Fréchet distributions such as the exponentiated-Fréchet, beta-Fréchet, Lomax–Fréchet, and Kumaraswamy Marshall–Olkin Fréchet.

**Keywords:** Cramér–von Mises estimation; engineering data; extreme value distribution; Fréchet distribution; maximum product of spacing estimators; simulations **Mathematics Subject Classification:** 60E05, 62F10

# 1. Introduction

Data analysis has been received a great interest in several applied fields such as medicine, reliability analysis, engineering, environmental studies, and economics. Several authors have proposed more flexible statistical distributions to model and predict various experimental and phenomenal data encountered in applied fields.

The Frchet (Fr) distribution is also known as the inverse-Weibull distribution and it is one of the useful distributions in extreme value theory. The Fr model has some important applications in life testing, floods, earthquakes, and wind speeds, among others. Further information about the applications of the Fr distribution can be explored in [1-5].

The Fr distribution is specified by the following cumulative distribution function (CDF)

$$F(x;\theta,\lambda) = e^{-\lambda x^{-\theta}}, \qquad \theta,\lambda > 0, \qquad x > 0.$$
(1.1)

Its probability density function (PDF) reduces to

$$f(x;\theta,\lambda) = \theta \lambda x^{-(\theta+1)} e^{-\lambda x^{-\theta}}, \qquad \theta,\lambda > 0, \qquad x > 0, \qquad (1.2)$$

where  $\theta$  and  $\lambda$  are, respectively, the shape and scale parameters. The PDF (1.2) exhibits a unimodal shape or a decreasing shape depending on  $\theta$  while its hazard rate function (HRF) is always unimodal.

There is a clear need to define and develop more flexible versions of the Fr model using the well-known families to model several datasets encountered in many applied fields such as medicine, geology, engineering, and economics, among others. Hence, many authors have proposed several generalized forms of the Fr distribution to improve its flexibility and capability in modeling real-life data. Some notable extensions are the following: the exponentiated-Fr [6], beta-Fr [7], Marshall–Olkin Fr [8], transmuted Marshall–Olkin Fr [9], Weibull–Fr [10], beta exponential-Fr [11], Burr-X Fr [12], odd Lindley–Fr [13], logarithmic-transformed Fr [14], and modified Kies–Fr distributions [15].

This article introduces a new flexible extension of the Fr distribution called the extended Weibull–Fréchet (EWFr) distribution, which provides more flexibility to model real-life data than other competing distributions. Then, the first motivation to this article is devoted to introducing the EWFr distribution as a new extension of the Fr distribution via the extended Weibull-G (EW-G) family [16]. The useful characteristics of the EWFr distribution can be summarized as follows: The EWFr distribution is a more flexible version for the Fr distribution, and it improves the fitting of real-life data; it produces more flexible kurtosis than the baseline Fr model. The HRF of the EWFr distribution can exhibit an upside-down bathtub shape, an increasing shape, and a decreasing shape. Its density function can exhibit a symmetrical shape, a unimodal shape, an asymmetrical shape, a J shape, and a reversed-J shape. Furthermore, the EWFr distribution can be adopted to model various data in the medicine and engineering sciences. This fact was illustrated by modeling two real datasets from both fields, showing its superiority fit over other competing distributions.

Another motivation to this article is to show how several classical estimators of the EWFr distribution perform for different parameter combinations and several sample sizes. Hence, the EWFr parameters are estimated using different estimation approaches including: the maximum product of spacings estimators (MPSEs), least-squares estimators (LSEs), the right-tail Anderson-Darling estimators (RADEs), the maximum likelihood estimators (MLEs), the weighted least-squares estimators (WLSEs), the percentiles estimators (PCEs), the Cramr–von Mises estimators (CRVMEs), and the Anderson–Darling estimators (ADEs). Extensive simulation results were introduced to explore the performance of these estimators. Furthermore, these estimators are compared using partial and overall ranks to determine the best method for estimation the parameters of the EWFr distribution.

The paper is organized in six sections as follows: Section 2 introduces the EWFr distribution and its related functions. The distribution properties are determined in Section 3. Section 4 presents some

classical estimators of the EWFr parameters. The simulation results for the classical methods are provided in the same section. Two real-life datasets are fitted using the EWFr distribution in Section 5. Some final remarks are presented in Section 6.

# 2. The EWFr distribution

The EWFr distribution is constructed based on the EW-G family [16] which is specified, for any baseline CDF  $G(x; \zeta)$ , by the CDF

$$F(x; \vartheta, \varphi, \zeta) = 1 - \left\{ 1 + \varphi \left[ \frac{G(x; \zeta)}{1 - G(x; \zeta)} \right]^{\vartheta} \right\}^{\frac{-1}{\varphi}}, \qquad \vartheta, \varphi > 0, \qquad x \in \mathfrak{R}.$$
(2.1)

The corresponding PDF of (2.1) takes the form

$$f(x; \vartheta, \varphi, \zeta) = \frac{\vartheta g(x; \zeta) G(x; \zeta)^{\vartheta-1}}{\left[1 - G(x; \zeta)\right]^{\vartheta+1}} \left\{ 1 + \varphi \left[\frac{G(x; \zeta)}{1 - G(x; \zeta)}\right]^{\vartheta} \right\}^{\frac{-1}{\varphi} - 1}.$$
(2.2)

where  $g(x; \zeta) = dG(x; \zeta)/dx$  refers to the baseline density with parameter vector  $\zeta$ .

To this end, by inserting Eq (1.1) in (2.1), the CDF of the EWFr model follows as

$$F(x; \boldsymbol{\eta}) = 1 - \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\frac{1}{\varphi}}, \qquad \vartheta, \, \varphi, \, \lambda, \, \theta > 0, \qquad x > 0, \tag{2.3}$$

where  $\boldsymbol{\eta} = (\vartheta, \varphi, \lambda, \theta)^{\mathsf{T}}$ . The PDF of the EWFr model reduces to

$$f(x; \boldsymbol{\eta}) = \vartheta \,\theta \,\lambda \,x^{-(\theta+1)} e^{\lambda x^{-\theta}} \left( e^{\lambda x^{-\theta}} - 1 \right)^{-(\vartheta+1)} \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-(\frac{1}{\varphi}+1)}, \tag{2.4}$$

where  $\vartheta$ ,  $\varphi$  and  $\theta$  are shape parameters whereas  $\lambda$  is a scale parameter.

#### 2.1. Some Useful Functions and Shapes

The survival function (SF) of the EWFr distribution is given as

$$S(x; \boldsymbol{\eta}) = \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\frac{1}{\varphi}}.$$

Long-term SF (LT-SF) is a useful feature in the modeling process, because a portion of the population may no longer be eligible to the event of interest with probability p (see [17, 18]).

The general form of the LT-SF is  $S_{LT}(x; p, \eta) = p + (1 - p)S(x; \eta)$ , where  $S(x; \eta)$  denotes the SF of any distribution and p denotes the probability of being cured. Hence, the PDF of the LT-SF can be derived as

$$f_{LT}(x; p, \boldsymbol{\eta}) = -\frac{\partial}{\partial x} S_{LT}(x; p, \boldsymbol{\eta}) = (1 - p)f(x; p, \boldsymbol{\eta}), p \in (0, 1).$$

Using Eq (2.4), the PDF of the LT-SF of the EWFr distribution takes the form

$$f_{LT}(x; p, \boldsymbol{\eta}) = \frac{\vartheta \theta \lambda (1-p) e^{\lambda x^{-\theta}}}{x^{\theta+1}} \left( e^{\lambda x^{-\theta}} - 1 \right)^{-(\vartheta+1)} \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-(\frac{1}{\varphi}+1)}$$

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The HRF of the EWFr distribution takes the form

$$h(x; \boldsymbol{\eta}) = \frac{\vartheta \, \theta \, \lambda \, x^{-(\theta+1)} \, e^{\lambda x^{-\theta}} \, \left( e^{\lambda x^{-\theta}} - 1 \right)^{-(\vartheta+1)}}{1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta}}.$$

Its reversed HRF has the form

$$r(x; \boldsymbol{\eta}) = \frac{\vartheta \, \theta \, \lambda \, x^{-(\theta+1)} e^{\lambda x^{-\theta}} \left( e^{\lambda x^{-\theta}} - 1 \right)^{-(\vartheta+1)} \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\left(\frac{1}{\varphi} + 1\right)}}{1 - \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\frac{1}{\varphi}}}.$$

The odd ratio of the EWFr model is derived as

$$O(x; \boldsymbol{\eta}) = \frac{F(x|\boldsymbol{\eta})}{S(x|\boldsymbol{\eta})} = \left\{1 + \varphi \left(e^{\lambda x^{-\theta}} - 1\right)^{-\theta}\right\}^{\frac{1}{\varphi}} - 1.$$

Figure 1 presents some possible shapes of the EWFr PDF for different values of its parameters. The EWFr PDF can be a symmetrical shape, a unimodal shape, an asymmetrical shape, a J shape, and a reversed-J shape. The hazard rate plots of the EWFr model are depicted in Figure 2. The EWFr HRF can be an increasing shape, a unimodal shape, and a decreasing shape.



**Figure 1.** Possible shapes of the EWFr PDF for several values of  $\vartheta$ ,  $\varphi$ ,  $\lambda$  and  $\theta$ .

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**Figure 2.** Possible shapes of the EWFr HRF for several values of  $\vartheta$ ,  $\varphi$ ,  $\lambda$  and  $\theta$ .

#### **3.** The EWFr characteristics

# 3.1. Quantile function and median

The quantile function (QF) of the EWFr distribution, say, Q(u), can be calculated by solving F(x) = p in (2.3) in terms of p. Then, the EWFr QF follows as

$$Q(p) = \lambda^{\frac{1}{\theta}} \left\{ \ln \left( \frac{\left[ (1-p)^{-\varphi} - 1 \right]^{\frac{1}{\theta}}}{\varphi^{\frac{1}{\theta}} + \left[ (1-p)^{-\varphi} - 1 \right]^{\frac{1}{\theta}}} \right) \right\}^{\frac{1}{\theta}}, 0 (3.1)$$

The median of the EWFr distribution follows by substituting p = 0.5 in Eq (3.1).

#### 3.2. Linear representation

A useful linear representation for the PDF of the EWFr model is provided based on [16]. Alizadeh et al. [16] introduced a simple representation for the density of the EW-G class as follows

$$f(x) = \sum_{w,u=0}^{\infty} \psi_{w,u} h_{\vartheta w+u}(x), \qquad (3.2)$$

where  $\psi_{w,u} = -\varphi^w \Gamma(\vartheta w + u) \left(-\frac{1}{\varphi}\right)_w / [w! \ u! \Gamma(\vartheta w)]$  and

$$h_{\vartheta w+u}(x) = (\vartheta w+u) g(x) G(x)^{\vartheta w+u-1},$$

is the exponentiated-G PDF with a power parameter ( $\vartheta w + u$ ) > 0. Using Eqs (1.1) and (1.2) of the Fr distribution, Eq (3.2) can be expressed as

$$f(x) = \sum_{w,u=0}^{\infty} \psi_{w,u} \left(\vartheta w + u\right) \theta \lambda x^{-(\theta+1)} e^{-(\vartheta w+u)\lambda x^{-\theta}}.$$
(3.3)

Equation (3.3) can be rewritten as

$$f(x) = \sum_{w,u=0}^{\infty} \psi_{w,u} \ g_{(\vartheta w+u)}(x),$$
(3.4)

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where  $g_{(\vartheta w+u)}(x)$  denotes the Fr PDF with two-parameter  $\theta$  and  $(\vartheta w + u)\lambda$ . Then, the density function of the EWFr model is expressed as a linear representation of Fr densities. Let *Y* be a random variable having the Fr distribution in (1.1). Hence, the *r*-th ordinary,  $\mu'_{r,Y}$ , and incomplete moments,  $\phi_{(r,Y)}(t)$ , of *Y* are, respectively, given (for  $r < \lambda$ ,) by

$$\mu_{r,Y}' = \lambda^{\frac{r}{\theta}} \Gamma(1 - \frac{r}{\theta}) \quad \text{and} \quad \phi_{(r,Y)}(t) = \lambda^{\frac{r}{\theta}} \gamma(1 - \frac{r}{\theta}, \lambda t^{-\theta}),$$

where  $\Gamma(s) = \int_0^\infty w^{s-1} e^{-w} dw$  is the complete gamma function (GF) and  $\gamma(s, z) = \int_0^z w^{s-1} e^{-w} dw$  is the lower incomplete GF.

#### 3.3. Moments and incomplete moments

This section was devoted to deriving the r-th ordinary moment and incomplete moments of the EWFr distribution.

**Proposition:** Based on (3.4), the *r*-th moment of the EWFr distribution is defined by

$$\mu'_{r} = \sum_{w,u=0}^{\infty} \psi_{w,u} \int_{0}^{\infty} x^{r} g_{\vartheta w+u}(x) dx \text{ for } r \in \mathbb{N}.$$
  
$$\mu'_{r} = \sum_{w,u=0}^{\infty} \psi_{w,u} \left(\vartheta w+u\right)^{\frac{r}{\theta}} \Gamma\left(1-\frac{r}{\theta}\right).$$
(3.5)

Setting r = 1 in Eq (3.5), we get the mean of x.

The *s*-th incomplete moment, say  $\phi_s(x)$ , of the EWFr distribution takes the form

$$\phi_{s}(t) = \int_{0}^{t} x^{s} f(x) dx = \sum_{w,u=0}^{\infty} \psi_{w,u} \int_{0}^{t} x^{s} g_{(\vartheta w+u)}(x) dx.$$

Then, we obtain  $(fors < \theta)$ 

$$\phi_s(t) = \sum_{w,u=0}^{\infty} \psi_{w,u} \, \left(\vartheta \, w + u\right)^{\frac{s}{\theta}} \, \gamma \left(1 - \frac{s}{\theta}, \, (\vartheta \, w + u) \lambda \, t^{-\theta}\right).$$

The mean ( $\mu$ ), variance ( $\sigma^2$ ), skewness ( $\xi_1(X)$ ), and kurtosis ( $\xi_2(X)$ ) of the EWFr distribution are calculated numerically with  $\lambda = 1$  and different values of  $\vartheta$ ,  $\varphi$  and  $\theta$ . Table 1 displays these numerical results. Table 1 shows that the EWFr model can be right-skewed and it can be leptokurtic (i.e.,  $\xi_2(X) > 3$ ).

**Table 1.** Some numerical values for  $\mu$ ,  $\sigma^2$ ,  $\xi_1(X)$ , and  $\xi_2(X)$  of the EWFr distribution with  $\lambda = 1$  and different values of  $\vartheta$ ,  $\varphi$  and  $\theta$ .

$\eta^{ op}$	$\mu$	$\sigma^2$	$\xi_1(X)$	$\xi_2(X)$
$(\vartheta = 1.50, \ \varphi = 0.50, \ \theta = 1.50)$	1.3511	0.4752	3.1421	56.2021
$(\vartheta = 2.50, \ \varphi = 0.75, \ \theta = 1.75)$	1.2701	0.1329	1.9395	16.1410
$(\vartheta = 2.25, \ \varphi = 1.00, \ \theta = 2.75)$	1.1904	0.0780	2.2648	19.2907
$(\vartheta = 2.25, \ \varphi = 0.25, \ \theta = 0.75)$	1.6154	0.7600	1.6132	9.0569
$(\vartheta = 2.25, \ \varphi = 1.25, \ \theta = 3.25)$	1.1831	0.0747	2.8068	27.8064
$(\vartheta = 5.50, \ \varphi = 1.50, \ \theta = 2.50)$	1.1927	0.0223	1.7507	11.7806

#### 3.4. Order statistics

Let  $X_1, X_2, ..., X_n$  be a random sample from the EWFr (2.4) and  $X_{1:n} \le X_{2:n} \le \cdots \le X_{n:n}$  be their corresponding order statistics (OS). The PDF and the CDF of the of *r*-th OS, say,  $X_{r:n}$  and  $1 \le r \le n$  are, respectively, defined by

$$f_{r:n}(x) = \frac{n!}{(n-r)! (r-1)!} [F(x)]^{r-1} [1 - F(x)]^{n-r} f(x)$$
  
$$= \frac{n!}{(n-r)! (r-1)!} \sum_{u=0}^{n-r} (-1)^u {\binom{n-r}{u}} [F(x)]^{r-1+u} f(x)$$
(3.6)

and (for k = 1, 2, ..., n)

$$F_{r:n}(x) = \sum_{l=k}^{n} \binom{n}{l} [F(x)]^{l} [1 - F(x)]^{n-l} = \sum_{l=k}^{n} \sum_{u=0}^{n-r} (-1)^{u} \binom{n}{l} \binom{n-r}{u} [F(x)]^{l+u}.$$
 (3.7)

Using Eqs (3.6) and (3.7), the PDF and CDF of the r-th OS of the EWFr reduce to

$$f_{r:n}(x) = \frac{\vartheta \ \theta \ \lambda \ x^{-(\theta+1) \ n!}}{(r-1)!(n-r)!} e^{\lambda t^{-\theta}} \left( e^{\lambda x^{-\theta}} - 1 \right)^{-(\vartheta+1)} \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-(\frac{1}{\varphi}+1)} \\ \sum_{u=0}^{n-r} (-1)^u \binom{n-r}{u} \left[ 1 - \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\frac{1}{\varphi}} \right]^{r+u-1}$$

and

$$F_{r:n}(x) = \sum_{l=k}^{n} \sum_{u=0}^{n-r} (-1)^u \binom{n}{l} \binom{n-r}{u} \left[ 1 - \left\{ 1 + \varphi \left( e^{\lambda x^{-\theta}} - 1 \right)^{-\theta} \right\}^{-\frac{1}{\varphi}} \right]^{l+u}.$$

### 4. Estimation and simulations

In this section, the EWFr parameters  $\vartheta$ ,  $\lambda$ ,  $\varphi$  and  $\theta$  are estimated using different frequentist approaches. We also provide detailed simulation results to compare and order their performances using partial and overall ranks.

#### 4.1. Maximum likelihood estimators

The MLEs of the parameters  $\vartheta$ ,  $\lambda$ ,  $\varphi$  and  $\theta$  of the EWFr distribution are introduced in this subsection. Let  $x_1, \ldots, x_n$  be a sample from the EWFr distribution in (2.4). Hence, the log-likelihood function of  $\eta = (\vartheta, \varphi, \lambda, \theta)^{\mathsf{T}}$  takes the form

$$\begin{split} l(\boldsymbol{\eta}; \boldsymbol{x}) &= n \log \left( \vartheta \right) + n \log \left( \theta \right) + n \log \left( \lambda \right) - \left( \theta + 1 \right) \sum_{i=1}^{n} \log(x_i) - \left( \vartheta + 1 \right) \sum_{i=1}^{n} \log \left( e^{\lambda x_i^{-\theta}} - 1 \right) \\ &+ \lambda \sum_{i=1}^{n} x_i^{-\theta} - \left( \frac{1}{\varphi} + 1 \right) \sum_{i=1}^{n} \log \left\{ 1 + \varphi \left( e^{\lambda x_i^{-\theta}} - 1 \right)^{-\vartheta} \right\}. \end{split}$$

The MLEs follow by maximizing the above equation by several programs such as SAS (PROC NLMIXED) or R (optim function).

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#### 4.2. Ordinary and weighted least-squares estimators

Let  $x_{(1)}, x_{(2)}, \dots, x_{(n)}$  be the OS of a random sample from the PDF (2.4), then the LSEs [19] of the EWFr parameters are obtained by minimizing the function:

$$S(\eta) = \sum_{i=1}^{n} \left[ F(x_{(i)}) - \frac{i}{n+1} \right]^2.$$

Similarly, these estimators are also obtained by solving the following equation (for k = 1, 2, 3, 4)

$$\sum_{i=1}^{n} \left( \left\{ 1 + \varphi \left( e^{\lambda x_i^{-\theta}} - 1 \right)^{-\vartheta} \right\}^{-\frac{1}{\varphi}} - \frac{i}{n+1} \right) \Omega_k \left( x_{(i)} | \boldsymbol{\eta} \right) = 0,$$

where

$$\Omega_{1}(x_{(i)}|\boldsymbol{\eta}) = \frac{\partial}{\partial \vartheta} F(x_{(i)}|\boldsymbol{\eta}) = \frac{1}{\varphi} \left(1 + \varphi w_{i}^{-\vartheta}\right)^{-(\frac{1}{\varphi}+1)} - \varphi w_{i}^{-\vartheta} \ln w_{i},$$
  

$$\Omega_{2}(x_{(i)}|\boldsymbol{\eta}) = \frac{\partial}{\partial \varphi} F(x_{(i)}|\boldsymbol{\eta}) = \left(1 + \varphi w_{i}^{-\vartheta}\right)^{-\frac{1}{\varphi}} \ln\left(1 + \varphi w_{i}^{-\vartheta}\right),$$
  

$$\Omega_{3}(x_{(i)}|\boldsymbol{\eta}) = \frac{\partial}{\partial \lambda} F(x_{(i)}|\boldsymbol{\eta}) = \frac{1}{\varphi} \left(1 + \varphi w_{i}^{-\vartheta}\right)^{-(\frac{1}{\varphi}+1)} - \vartheta \varphi x_{i}^{-\theta} w_{i}^{-(\vartheta+1)} e^{\lambda x_{i}^{-\theta}}$$

and

$$\Omega_4\left(x_{(i)}|\boldsymbol{\eta}\right) = \frac{\partial}{\partial\theta} F\left(x_{(i)}|\boldsymbol{\eta}\right) = \vartheta \lambda w_i^{-(\vartheta+1)} e^{\lambda x_i^{-\theta}} x_i^{-\theta} \left(1 + \varphi w_i^{-\vartheta}\right)^{-(\frac{1}{\varphi}+1)} \ln x_i,$$

where  $w_i = e^{\lambda x_i^{-\theta}} - 1$ . The solution of  $\Omega_k$  for k = 1, 2, 3, 4 may be obtained numerically.

The WLSEs of the EWFr parameters can be determined by minimizing the equation (see [19]):

$$W(\boldsymbol{\eta}) = \sum_{i=1}^{n} \frac{(n+1)^2 (n+2)}{i (n-i+1)} \left[ F(x_{(i)} | \boldsymbol{\eta}) - \frac{i}{n+1} \right]^2.$$

# 4.3. Maximum product of spacing and Cramér-von Mises estimators

The uniform spacings of a random sample from the EWFr distribution are defined (for i = 1, 2, ..., n + 1) by  $D_i(\eta) = F(x_{(i)}|\eta) - F(x_{(i-1)}|\eta)$ , where  $F(x_{(0)}|\eta) = 0$ ,  $F(x_{(n+1)}|\eta) = 1$  and  $\sum_{i=1}^{n+1} D_i(\eta) = 1$ . The MPSEs of the EWFr parameters can be determined by maximizing the following geometric mean (GM) of spacings

$$G(\boldsymbol{\eta}) = \left[\prod_{i=1}^{n+1} D_i(\boldsymbol{\eta})\right]^{\frac{1}{n+1}}$$

or by maximizing the logarithm of the GM of sample spacings

$$H(\boldsymbol{\eta}) = \frac{1}{n+1} \sum_{i=1}^{n+1} \log D_i(\boldsymbol{\eta}),$$

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The CRVMEs can be obtained using the difference between the estimated and empirical CDFs. The CRVMEs [20] of the EWFr parameters are determined by minimizing

$$C(\boldsymbol{\eta}) = \frac{1}{12n} + \sum_{i=1}^{n} \left[ F(x_{(i)}|\boldsymbol{\eta}) - \frac{2i-1}{2n} \right]^{2}.$$

### 4.4. Anderson–Darling and right-tail Anderson–Darling estimators

The ADEs [21] of the EWFr parameters are calculated by minimizing

$$A(\eta) = -n - \frac{1}{n} \sum_{i=1}^{n} (2i - 1) \left[ \log F(x_{(i)}|\eta) + \log S(x_{(i)}|\eta) \right].$$

The RADEs of the EWFr parameters are calculated by minimizing

$$R(\boldsymbol{\eta}) = \frac{n}{2} - 2\sum_{i=1}^{n} F(x_{i:n}|\boldsymbol{\eta}) - \frac{1}{n}\sum_{i=1}^{n} (2i-1)\log S(x_{n+1-i:n}|\boldsymbol{\eta})$$

#### 4.5. Percentile estimators

Consider the unbiased estimator of  $F(x_{(i)}|\boldsymbol{\eta})$  which is defined by  $u_i = i/(n+1)$ . Hence, the PCEs of the EWFr parameters can be calculated by minimizing

$$P(\boldsymbol{\eta}) = \sum_{i=1}^{n} \left( x_{(i)} - \left\{ \ln \left( \frac{[(1-u_i)^{-\varphi} - 1]^{\frac{1}{\theta}}}{\varphi^{\frac{1}{\theta}} + [(1-u_i)^{-\varphi} - 1]^{\frac{1}{\theta}}} \right)^{\lambda} \right\}^{\frac{1}{\theta}} \right)^{2}.$$

#### 4.6. Simulation analysis

To compare and explore the behavior of different estimators of the EWFr parameters, we presented the numerical simulation results and ranked them with respect to their: average of absolute biases  $(|Bias(\widehat{\eta})|), |Bias(\widehat{\eta})| = \frac{1}{N} \sum_{i=1}^{N} |\widehat{\eta} - \eta|$ , average of mean relative errors (MREs),  $MREs = \frac{1}{N} \sum_{i=1}^{N} |\widehat{\eta} - \eta|/\eta$ , and average mean square errors (MSEs),  $MSEs = \frac{1}{N} \sum_{i=1}^{N} (\widehat{\eta} - \eta)^2$ .

The following algorithm can be adopted to explore the behavior of different estimators of the EWFr parameters:

**Step 1:** A random sample  $x_1, x_2, ..., x_n$  of sizes n = 20, 80, 200, and 500 are generated from the QF (3.1).

**Step 2:** The required results are obtained based on eight combinations of the parameters  $\vartheta = \{0.25, 0.75, 1.75, 3.50, 3.50\}, \varphi = \{0.50, 0.75, 2.00, 2.50, 3.50\}, \lambda = \{0.25, 0.50, 1.50, 3.00, 4.25\}$  and  $\theta = \{0.25, 1.25, 2.50\}.$ 

**Step 3:** Each sample is replicated N = 5,000 times.

**Step 4:** Results of the biases, MSEs, and MREs are computed for the eight combinations, and to save more space, we present just the result of 5 combinations in Tables 2–6.

All computations are obtained using **R** software (version 4.0.2) [22].

**Table 2.** Simulation results for  $\boldsymbol{\eta} = (\vartheta = 1.75, \varphi = 0.5, \lambda = 0.25, \theta = 1.25)^{\mathsf{T}}$ .

n	Est.	Est. Par.	MLEs	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADEs
		$\hat{artheta}$	0.58150 {3}	$0.59477^{4}$	0.53422 {1}	0.54541 23	0.60555 <sup>{5}</sup>	0.71685 <sup>{8}</sup>	0.67445 77	0.61466 {6}
	BIAS	$\hat{arphi}$	0.47550 44	0.54812 7	0.46820 23	0.53093 6	0.19120 11	0.55104 88	0.46906 {3}	0.51894 [5]
	DIT IS	λ	0.08930 23	0.10351 <sup>{7}</sup>	0.08402 {1}	0.10057 5	0.10077 <sup>{6}</sup>	0.11204 [8]	0.09758 {3}	0.09866 <sup>{4}</sup>
		$\hat{ heta}$	0.49749 <sup>{7}</sup>	0.47497 <sup>{4}</sup>	0.46884 {3}	0.48364 {5}	0.45238 {1}	0.53285 <sup>{8}</sup>	0.46505 <sup>{2}</sup>	$0.48608^{\{6\}}$
		$\hat{artheta}$	0.47425 {2}	0.52401 44	0.52337 {3}	0.46463 {1}	0.54588 [6]	0.64689 {8}	0.59726 <sup>{7}</sup>	0.53428 {5}
	MSEs	$\hat{arphi}$	0.27624 {1}	0.38734 <sup>{7}</sup>	0.29294 <sup>{2}</sup>	0.36212 6	0.31849 4	0.39577 <sup>{8}</sup>	0.30160 {3}	0.35513 {5}
20	1010125	Â	0.01107 <sup>{2}</sup>	0.01423 <sup>{7}</sup>	0.00983 {1}	0.01356 {5}	0.01357 [6]	0.01651 {8}	0.01289 {3}	0.01296 4
		$\hat{\theta}$	0.37764 {7}	0.33320 {3}	0.35778 44	0.36101 (5)	0.29381 {1}	0.38887 [8]	0.32193 {2}	0.36311 [6]
		θ	0.33229 [3]	0.33987 4	0.17807 {1}	0.31166 {2}	0.34603 (5)	0.40963 [8]	0.38540 {/}	0.35123 [6]
	MREs	$\hat{arphi}$	0.95101 [3]	1.09624 {7}	0.93641 {1}	1.06186 [6]	0.97231 44	1.10208 [8]	0.93811 {2}	1.03788 {5}
		$\hat{\lambda}$	0.35719 {2}	0.41405 {/}	0.33607 {1}	0.40230 {5}	0.40307 [6]	0.44817 [8]	0.39034 {3}	0.39464 4
		θ	0.39799 17	0.37998 (4)	0.37507 (3)	0.38691 (5)	0.36190 {1}	0.42628 (8)	0.37204 <sup>{2</sup> }	0.38886
		$\sum_{n \in \mathbb{R}} Ranks$	43 {2}	65 {/}	23 {1}	53 (5)	46 (4)	96 [8]	44 {5}	62 (6)
		θ	0.65720 {2}	0.73260 {0}	0.64523 {1}	0.71780 (4)	0.67014 (3)	0.78730 [8]	0.72979 (3)	0.78285 {/}
	BIAS	$\hat{\varphi}$	0.29918 (4)	0.36951 {/}	0.27270 {1}	0.36572 (6)	0.28461 {2}	0.43485 [8]	0.29440 {3}	0.31911 (5)
	1	$\hat{\lambda}$	0.06946 (2)	0.08089 17	0.06755 (1)	0.07955 (0)	0.07162 (4)	0.07099 (3)	0.07301 (3)	0.08347 [8]
		$\hat{\theta}$	0.41705 (0)	0.41605 (5)	0.41277 (4)	0.43216 {/}	0.34179 {1}	0.38656	0.37952 {2}	0.45183 (8)
		θ	0.53699 (1)	0.65792 {3}	0.71214 {/}	0.63722 (3)	0.62257 {2}	0.71659 [8]	0.64185 (4)	0.70948 (0)
00	MSEs	$\hat{\varphi}$	0.12478 (3)	0.19502 {/}	0.11097 {1}	0.18881 (0)	0.12320 <sup>{2}</sup>	0.28084 [8]	0.13150 (4)	0.15268 (3)
80		λ â	$0.006/2^{(1)}$	0.00884	0.00693 (2)	0.00853	$0.00712^{(4)}$	$0.00711^{(3)}$	$0.00^{7}/40^{-13}$	0.00932
		$\theta$	0.26049 (3)	0.25434 (4)	0.30456	0.27845	0.16695	0.24032 (3)	0.20942 (2)	0.29980 (7)
		θ ^	0.37554 (2)	0.41863	0.21508 (1)	0.41017 (4)	0.38294	0.44989 187	0.417/02 (3)	0.44734 17
	MREs	$\hat{\varphi}$	0.59835	0.73902	0.54540	0.73144	0.56922	0.86970	0.58881 (5)	0.63822
		л â	0.27786 (2)	0.32356 (7)	0.27019 (1)	0.31820 (0)	0.28647	0.28396 <sup>(3)</sup>	$0.29205^{(3)}$	0.33388 [8]
		$\theta$ $\sum D = 1 + 1 + 1 + 1 + 1 + 1 + 1 + 1 + 1 + 1$	0.33364	0.33284	0.33022	$0.34573^{(1)}$	0.27343	0.30924	$0.30361^{(2)}$	0.36147 (8)
		<u>}</u> Kanks	0 54244 {1}	0 74060 [6]	0 66029 {4}	0.72065 {5}	0 56860 {2}	0.78528 {8}	$\frac{43}{0.64417}$	0.74318 {7}
		ô	0.34244 $0.20343^{3}$	0.77000	0.18758 {1}	0.72003	$0.19120^{\{2\}}$	$0.33648^{\{8\}}$	0.04417 $0.21264^{\{4\}}$	0.74910
	BIAS	ĵ	0.05329 {1}	0.27714 0.07033 $\{7\}$	0.05985 <sup>{3}</sup>	0.06898 <sup>{6}</sup>	0.15120	0.05992 <sup>{4</sup> }	0.06059 <sup>{5}</sup>	0.21770
		Â	$0.30910^{\{2\}}$	0.38666 <sup>{6}</sup>	0.37834 <sup>{5}</sup>	$0.39240^{\{7\}}$	$0.26754^{\{1\}}$	0.34697 <sup>{4}</sup>	$0.32167^{\{3\}}$	$0.42492^{\{8\}}$
		พิ	0.39682 <sup>{1}</sup>	0.65105 <sup>{6}</sup>	$0.72704^{\{8\}}$	$0.62292^{\{4\}}$	$0.51529^{\{2\}}$	$0.70871^{7}$	0.53035 <sup>{3}</sup>	0.65051 {5}
		ŵ	0.06226 <sup>[3]</sup>	0.11401 <sup>{7}</sup>	0.05424 {1}	$0.11214^{\{6\}}$	$0.05643^{\{2\}}$	0.18593 <sup>{8}</sup>	$0.06927^{\{4\}}$	$0.07447^{\{5\}}$
200	MSEs	λ	0.00405 {1}	0.00665 {7}	0.00571 {5}	0.00644 {6}	0.00452 {2}	0.00513 {4}	0.00507 {3}	0.00753 [8]
		$\hat{ heta}$	0.14395 {2}	0.20851 {5}	0.26568 {8}	0.22075 [6]	0.10322 {1}	0.19203 {4}	0.14560 {3}	0.25309 {7}
		$\hat{\vartheta}$	0.30997 {2}	0.42320 [6]	0.22010 {1}	0.41180 {5}	0.32491 {3}	0.44873 {8}	0.36810 {4}	0.42468 {7}
	) (DE	$\hat{arphi}$	0.40685 {3}	0.55429 {7}	0.37516 {1}	0.55256 863	0.38240 {2}	0.67295 {8}	0.42528 4	0.43980 {5}
	MRES	λ	0.21316 {1}	0.28134 7	0.23940 {3}	0.27593 863	0.22675 {2}	0.23969 43	0.24238 (5)	0.30470 [8]
		$\hat{ heta}$	0.24728 {2}	0.30933 863	0.30267 {5}	0.31392 7	0.21403 11	0.27758 43	0.25734 {3}	0.33994 88
		$\sum Ranks$	22 {1.5}	77 {7}	45 {4}	70 {5}	22 {1.5}	71 {6}	44 {3}	81 {8}
		$\overline{\hat{\vartheta}}$	0.40816 {2}	0.63822 [6]	0.59712 43	0.62303 {5}	0.38723 {1}	0.72351 [8]	0.50869 {3}	0.64537 77
	DIACI	$\hat{arphi}$	0.13392 33	0.19775 863	0.12778 {1}	0.19966 {7}	0.13208 23	0.26446 {8}	0.14927 43	0.15660 {5}
	DIA5	λ	0.03902 [1]	0.05985 7	$0.04987$ <sup>{4}</sup>	0.05913 863	0.04014 23	0.05408 [5]	0.04806 {3}	0.06618 [8]
		$\hat{ heta}$	0.21255 2	0.32416 463	0.31285 (5)	0.33146 7	0.18308 11	0.30762 43	0.24975 33	0.35803 [8]
		$\hat{artheta}$	0.24842 {1}	0.52025 (5)	0.58302 7	0.50041 43	0.33081 23	0.63280 {8}	0.37071 {3}	0.52827 [6]
	MSE	$\hat{arphi}$	0.02762 {3}	0.05937 <sup>{6}</sup>	0.02573 {1}	0.06052 7	0.02727 <sup>{2}</sup>	0.12045 [8]	0.03399 43	0.03724 {5}
500	MISES	λ	0.00224 {1}	0.00478 7	0.00422 {5}	0.00470 863	0.00257 23	0.00402 {4}	0.00325 {3}	0.00569 88
		$\hat{ heta}$	$0.06777^{\{2\}}$	0.14391 <sup>{5}</sup>	0.19157 <sup>{8}</sup>	0.15413 68	0.05743 {1}	0.14128 44	0.08776 {3}	0.17743 77
		$\hat{artheta}$	0.23323 {3}	0.36470 863	$0.19904^{\{1\}}$	0.35602 {5}	0.22127 {2}	0.41343 88	0.29068 44	0.36879 7
	MRFs	$\hat{arphi}$	0.26784 {3}	0.39550 863	0.25556 {1}	0.39932 <sup>{7}</sup>	0.26417 {2}	0.52891 {8}	0.29853 44	0.31320 (5)
		λ	0.15609 {1}	0.23939 <sup>{7}</sup>	0.19949 <sup>{4}</sup>	0.23651 [6]	0.16056 <sup>[2]</sup>	0.21632 {5}	0.19222 [3]	0.26474 [8]
		$\theta$	0.17004 <sup>{2}</sup>	0.25933 [6]	0.25028 {5}	0.26517 <sup>{7}</sup>	0.14646 {1}	0.24610 44	0.19980 {3}	0.28642 [8]
		$\sum Ranks$	24 {2}	73 {5.5}	46 44	73 {5.5}	$20^{\{1\}}$	74 <sup>{7}</sup>	40 {3}	82 {8}

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**Table 3.** Simulation results for  $\boldsymbol{\eta} = (\vartheta = 1.75, \varphi = 2, \lambda = 1.5, \theta = 2.5)^{\mathsf{T}}$ .

n	Est.	Est. Par.	MLEs	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADEs
		$\hat{\vartheta}$	0.58249 {3}	0.53932 {1}	0.56241 {2}	0.61397 {5}	0.60739 {4}	0.82341 {8}	0.68123 {7}	0.66592 {6}
		ŵ	0.70201 {2}	0.74992 {4}	0.71005 {3}	0.99334 {7}	0.41044 {1}	1.11769 {8}	0.88784 {5}	0.97457 [6]
	BIAS	λ	0.26246 {1}	0.28717 {3}	0.28400 {2}	0.42881 {5}	0.40138 {4}	0.64545 <sup>{8}</sup>	0.43589 {7}	0.43050 [6]
		Â	0.83485 {6}	0.91523 {7}	0.92408 {8}	0.76222 {3}	0.77841 {4}	0.78092 {5}	0.74535 {1}	0.74828 {2}
		ŵ	0.45340 {3}	$0.42087^{\{1\}}$	$0.44427^{\{2\}}$	0.51220 <sup>{4}</sup>	0.51720 {5}	0.77654 <sup>{8}</sup>	0.58720 {7}	0.56968 <sup>{6}</sup>
		ŵ	0.65182 {1}	0.75540 {3}	0.66110 <sup>{2}</sup>	1.25282 {7}	0.95641 <sup>{4}</sup>	1.49737 <sup>{8}</sup>	1.01553 <sup>{5}</sup>	1.19476 <sup>{6}</sup>
20	MSEs	ĵ	0.07059 {1}	0.08751 {3}	0.08543 {2}	0.29909 {7}	$0.24742^{\{4\}}$	0.52940 {8}	0.29252 [6]	0.29154 {5}
		Â	$0.84612^{\{6\}}$	1.00860 <sup>{8}</sup>	$1.00532^{\{7\}}$	0.71906 {3}	0.75325 <sup>{4}</sup>	0.82231 {5}	0.69565 {1}	0.70768 {2}
		ŵ	0.33285 <sup>{3}</sup>	0.30818 {1}	$0.32138^{\{2\}}$	0.35084 {5}	0.34708 [4]	0.47052 [8]	0.38927 <sup>{7}</sup>	0.38052 <sup>{6}</sup>
		ŵ	0.35100 {1}	0.37496 {3}	0.35503 {2}	0.49667 <sup>{7}</sup>	0.42863 <sup>{4}</sup>	0.55884 {8}	0.44392 {5}	0.48728 [6]
	MREs	ĵ	0.17497 {1}	0.19145 {3}	0.18933 {2}	0.28587 {5}	0.26758 <sup>{4}</sup>	0.43030 <sup>{8}</sup>	0.29059 {7}	0.28700 {6}
		Â	$0.33394^{\{6\}}$	0.36609 <sup>{7}</sup>	0.36963 <sup>{8}</sup>	0 30489 <sup>{3}</sup>	0.31136 <sup>{4}</sup>	0.31237 <sup>{5}</sup>	$0.29814^{\{1\}}$	$0.29931^{\{2\}}$
		$\Sigma Ranks$	<b>34</b> <sup>{1</sup> }	<b>44</b> <sup>{3}</sup>	<b>42</b> <sup>{2}</sup>	<b>61</b> <sup>{7}</sup>	<b>46</b> <sup>{4}</sup>	<b>87</b> <sup>{8}</sup>	<b>59</b> {5.5}	<b>59</b> {5.5}
		<u>n</u> n	0 49925 {2}	0.47135 {1}	0.52766 {3}	0.66329 {5}	0 55966 {4}	0 89224 [8]	0.67228 [6]	0 69447 {7}
		ô	0.52063 <sup>{1}</sup>	0.55025 {3}	$0.52909^{\{2\}}$	0.70683 <sup>{7}</sup>	0.56354 <sup>{4}</sup>	0.93582 <sup>{8}</sup>	0.63998 <sup>{5}</sup>	0.69954 <sup>{6}</sup>
	BIAS	ĵ	0.25198 {1}	0.25883 {3}	0.32505	0.43311 {7}	0 30034 <sup>{4</sup> }	0.66563 <sup>{8}</sup>	0.40607 <sup>{5}</sup>	0.43206 [6]
		Â	0.23190 0.78480 $\{6\}$	0.23003 $0.82532$ <sup>{7}</sup>	0.83697 <sup>{8}</sup>	0.69063 <sup>{3}</sup>	0.58559 {1}	$0.72122^{5}$	$0.63842^{\{2\}}$	$0.69164^{\{4\}}$
		พิ	$0.31719^{\{2\}}$	0.30026 {1}	$0.34946^{\{3\}}$	0.55815 <sup>{5}</sup>	0.49857 <sup>{4}</sup>	0.86629 {8}	0 56483 [6]	0 59417 <sup>{7}</sup>
		ŵ	$0.27881^{\{1\}}$	$0.32553^{\{3\}}$	0.34940 0.29109 <sup>{2}</sup>	0.66883 <sup>{7}</sup>	0.44978 <sup>{4</sup> }	1.03760 <sup>{8}</sup>	0.55511 <sup>{5</sup> }	0.64550 <sup>{6}</sup>
80	MSEs	φ ĵ	0.06369 {1}	0.06795 <sup>{3}</sup>	0.29109	$0.29117^{\{7\}}$	$0.15725^{\{4\}}$	$0.54984^{\{8\}}$	$0.25250^{5}$	0.04330
00		Â	$0.66344^{5}$	0.73666 <sup>{7}</sup>	0.74889 {8}	0.29117 0.58730 <sup>{3}</sup>	0.15725 $0.45144^{\{1\}}$	0.66717 <sup>{6</sup> }	0.23230 0.51377 <sup>{2}</sup>	0.59201 <sup>[4]</sup>
		â	0.28529 {2}	0.75000	$0.7400^{-3}$	0.37902 <sup>{5</sup> }	0.31081 {4}	0.50985 {8}	$0.38416^{\{6\}}$	0.39684 <sup>[7]</sup>
		ô	0.26032 {1}	0.20934	0.30132 0.26454 $\{2\}$	0.35341 {7}	0.28177 {4}	0.30983	0.31999 {5}	0.34977 [6]
	MREs	φ ĵ	0.16798 {1}	0.17255 {3}	$0.20494^{\{2\}}$	0.28874 {7}	0.20023 [4]	0.40791	$0.31999^{-1}$	0.28804 [6]
		λ Δ	0.31302 [6]	0.33013 {7}	$0.17094^{[8]}$	$0.28874^{10}$	$0.20023^{\{1\}}$	0.28840 {5}	0.25537 <sup>[2]</sup>	0.27666 <sup>[4]</sup>
		$\sum Ranks$	0.51592 <b>20</b> {1}	0.55015 12 <sup>[3]</sup>	0.33479 45 <sup>{4}</sup>	66 <sup>{6</sup>	<b>30</b> <sup>{2</sup> }	88 {8}	54 <sup>{5</sup> }	60 <sup>{7</sup> }
		<u>n</u> n	0.47555 {3}	0 44985 {2}	0 51341 <sup>{4}</sup>	0.67707 <sup>{6}</sup>	0 42436 {1}	0.89805 <sup>{8}</sup>	0 66134 {5}	0.69895 <sup>{7}</sup>
		ô	0.47333	0.50837 <sup>{4}</sup>	0.51341 0.50279 <sup>{3}</sup>	0.56296 {7}	0.41044 {1}	0.81679 <sup>{8}</sup>	0.51451 <sup>{5}</sup>	0.55936 [6]
	BIAS	ĵ	$0.25019^{\{2\}}$	$0.25187^{\{4\}}$	$0.25076^{\{3\}}$	0.30270 0.41917 <sup>{7}</sup>	$0.21499^{\{1\}}$	$0.65544^{\{8\}}$	0.37297 <sup>{5}</sup>	$0.41435^{\{6\}}$
		Â	0.79159 {6}	0.23107 0.80107 <sup>{7}</sup>	0.23070 $0.82123^{\{8\}}$	$0.64647^{\{3\}}$	0.21122	0.75308 <sup>{5}</sup>	0.57297	0.66234 <sup>{4}</sup>
		พิ	0.75135	$0.23847^{\{1\}}$	0.02123	0.56097 <sup>{6}</sup>	$0.37531^{\{4\}}$	0.86783 <sup>{8}</sup>	$0.54545^{5}$	0.59107 <sup>{7}</sup>
		ŵ	0.25000	0.250 + 7 0.26082 <sup>{4}</sup>	$0.25349^{\{2\}}$	0.43952 {7}	0.25968 <sup>{3}</sup>	0.80523 {8}	0.37160 <sup>{5}</sup>	0.33107 0.43084 $\{6\}$
200	MSEs	φ ĵ	0.25154	0.06361 <sup>{3}</sup>	0.2537	0.25891 {7}	0.09970 {4}	0.52619 {8}	0.20829 {5}	0.45004
200		Â	0.64670 <sup>{5}</sup>	$0.66161^{\{6\}}$	$0.69240^{\{8\}}$	0.51309 {3}	$0.07577^{\{1\}}$	0.68347 {7}	$0.2002^{j}$ $0.41304^{\{2\}}$	0.53268 <sup>[4]</sup>
		พิ	$0.27175^{\{3\}}$	$0.25705^{\{2\}}$	0.29338 {4}	0.38690 {6}	0 24249 {1}	0.51317 <sup>{8}</sup>	0.37791 <sup>{5</sup> }	0.39940 {7}
		ŵ	0.27173 0.25057 <sup>{2}</sup>	$0.25419^{\{4\}}$	0.25139 {3}	$0.28148^{7}$	0.20522 {1}	0.40839 {8}	0.25725 (5)	0.27968 [6]
	MREs	ĵ	0.25057 0.16679 <sup>{2}</sup>	0.16791 <sup>{4}</sup>	0.25137 $0.16717^{\{3\}}$	0.20140	0.14333 {1}	0.43696 <sup>{8}</sup>	0.23725	0.27500
		Â	$0.31664^{\{6\}}$	$0.32043^{7}$	$0.32849^{\{8\}}$	0.25859 {3}	$0.17211^{\{1\}}$	$0.30123^{5}$	$0.27872^{\{2\}}$	0.27023 0.26494 <sup>{4}</sup>
		$\Sigma Ranks$	<b>35</b> <sup>{2}</sup>	<b>48</b> <sup>{3</sup> }	51 <sup>{4.5</sup> }	<b>69</b> {6.5}	20 {1}	<b>89</b> <sup>{8</sup> }	51 <sup>{4.5</sup> }	<b>69</b> {6.5}
		n nanks ŷ	$0.47314^{\{3\}}$	0 44590 {2}	0 51215 <sup>{4}</sup>	0 64549 {6}	0 23589 {1}	0 89642 {8}	0 57765 <sup>{5}</sup>	0.66728 <sup>{7}</sup>
		ô	0.50000 {5}	0.50031 <sup>{7}</sup>	0.50001 {6}	0.44580 <sup>{4}</sup>	$0.25497^{\{1\}}$	0.65012	0 39329 {2}	0.00720
	BIAS	ĵ	$0.25000^{\{2\}}$	0.25003 <sup>{4}</sup>	$0.25001^{\{3\}}$	0.38068 {6}	0.11500 {1}	0.63857 <sup>{8}</sup>	0.31593 {5}	0.38454 <sup>{7}</sup>
		Â	$0.79445^{\{5\}}$	$0.80134^{\{6\}}$	$0.82604^{8}$	$0.59243^{\{3\}}$	$0.25280^{\{1\}}$	0.80505 <sup>{7}</sup>	0 49356 <sup>{2}</sup>	$0.61821^{\{4\}}$
		ŵ	$0.23703^{\{3\}}$	$0.21187^{\{2\}}$	$0.27432^{\{4\}}$	0.51809 {6}	0 19898 {1}	0.86166 <sup>{8}</sup>	0 44538 {5}	0 54549 <sup>{7}</sup>
		ŵ	$0.25700^{\{3\}}$	$0.25034^{5}$	0.27132	$0.28548^{\{7\}}$	0.11705 {1}	0.56366 <sup>{8}</sup>	$0.22462^{\{2\}}$	0.25796 <sup>{6}</sup>
500	MSEs	ĵ	0.06250 {2.5}	$0.06252^{\{4\}}$	0.06250 {2.5}	0.20888 {6}	0.03935 {1}	0.50179 {8}	$0.14873^{5}$	$0.20922^{7}$
200		Â	0.63903 (5)	0.64949 {6}	0.68943 <sup>{7}</sup>	0.43597 {3}	0.12546 {1}	0.74499 {8}	$0.31912^{\{2\}}$	0.46804 {4}
		พิ	$0.27037^{\{3\}}$	$0.25480^{\{2\}}$	0 29266 <sup>{4}</sup>	0 36885 [6]	0 13479 {1}	0 51224 [8]	0 33008 {5}	0 38130 <sup>[7]</sup>
		ô	0 25000 {5.5}	0.25015 {7}	0 25000 {5.5}	0 22290 [4]	0 12748 {1}	0 33070 {8}	0 19665 <sup>[2]</sup>	0 21134 [3]
	MREs	ĵ	0 16667 {2.5}	0.16669 <sup>{4}</sup>	0 16667 {2.5}	0 25379 [6]	0.07667 {1}	0 42571 {8}	0 21062 {5}	0.25636 <sup>{7</sup> }
		Â	0.31778 {5}	0.32053 [6]	0.33042 [8]	$0.23697^{\{3\}}$	$0.10112^{\{1\}}$	0.32202 {7}	$0.19742^{\{2\}}$	$0.24728^{\{4\}}$
		$\sum Ranks$	<b>44.5</b> <sup>{3}</sup>	55 {4}	<b>58.5</b> <sup>{5}</sup>	<b>60</b> <sup>{6}</sup>	12 {1}	<b>94</b> <sup>{8}</sup>	<b>42</b> <sup>{2}</sup>	<b>66</b> <sup>{7}</sup>

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**Table 4.** Simulation results for  $\boldsymbol{\eta} = (\vartheta = 3.5, \varphi = 0.75, \lambda = 0.5, \theta = 1.25)^{\mathsf{T}}$ .

n	Est.	Est. Par.	MLEs	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADEs
		â	1 05031 (8)	0 (2005 (3)	0.0000 (6)	0 (1070 (2)	0.00700 (1)	1 000(0 (7)	0.02104 (5)	0 74051 (4)
		ϑ ^	1.05821 (6)	0.62985	0.96882 (3)	0.618/9 <sup>[2]</sup>	0.39728	1.00062 (7)	0.93184	0.74851 (4)
	BIAS	$\varphi$	0.58239	$0.61540^{(6)}$	$0.5/588^{(3)}$	$0.61442^{(7)}$	0.16810 <sup>(1)</sup>	0.59648 (8)	$0.54321^{(2)}$	$0.58344^{(3)}$
		л â	0.09205 [8]	$0.08984^{(3)}$	0.08/49 [5]	$0.09101^{(6)}$	0.07490 (1)	0.09181 (7)	0.08300 (2)	0.08890
		θ ô	0.65193 (8)	0.48317 (2)	0.52978 (6)	0.51841 (3)	0.37874	0.57738 (7)	0.51808 (5)	0.53685
		v ^	1.62863 [3]	0.80009 [2]	$1.26/47^{(6)}$	$0.83422^{(3)}$	$0.44242^{(1)}$	1.52407 (7)	1.26323	1.02485 [4]
20	MSEs	$\varphi_{\hat{\lambda}}$	0.42299 (5)	$0.4/351^{(5)}$	$0.43072^{(1)}$	$0.47310^{(7)}$	0.39842	0.45399 (8)	0.39008	0.43912 (3)
20		л â	0.01191 (8)	$0.01157^{(3)}$	0.01104 (3)	0.01199 (5)	0.00839	0.01226 (3)	0.01037 (2)	0.01130
		θ ô	0.36633 (8)	0.35145 (2)	0.39993	0.40530	0.22597	0.46238 (7)	0.39998	$0.42492^{(6)}$
		v ^	$0.30235^{(6)}$	0.1/996 (8)	$0.27681^{(3)}$	$0.1/680^{(2)}$	0.11351 (1)	0.28589	0.26624	0.21386
	MREs	$\varphi_{\hat{\lambda}}$	0.17652 (4)	0.82054 (5)	0.76784 [3]	0.81923	0.73001 (2)	0.19531 (8)	0.12428 (1)	0.17792 <sup>[3]</sup>
		л â	$0.18410^{(8)}$	$0.1/968^{(3)}$	0.17498 [5]	$0.18321^{(6)}$	0.14981 (1)	0.18363 (7)	$0.16/20^{(2)}$	0.17780 [4]
		$\theta$	0.52155	0.38654	0.42382 (3)	0.414/3 (4)	0.30299	0.46190 (*)	0.41446	0.42948 (6)
		$\sum_{n} Ranks$	81 17	53 (*)	50 (5)	60 <sup>[0]</sup>	14	82 (0)	35 (2)	57 (5)
		ϑ ^	1.30458	0.99944	1.19014	1.02885	0.28553	1.106/0	1.14568	1.11457
	BIAS	$\hat{\varphi}$	0.32983 (3)	0.40620 (0)	0.34267 (4)	0.40/46 (7)	0.27616 (1)	0.42471	0.32604 (2)	0.35934
		λ â	0.06496 (7)	0.06255	0.05976 (5)	0.06558 [8]	0.03810 (1)	0.06030 (4)	0.05724 (2)	0.06303 (0)
		θ	0.60157	0.45112 (2)	0.47825 (3)	0.49541	0.19099	0.45370 (3)	0.45453 (4)	0.50033
		θ	1.98067 [8]	1.42246 (2)	1.62/21 {/}	1.51028 (5)	0.28600 {1}	1.54773 (3)	1.54350 (4)	1.62460 (6)
0.0	MSEs	$\hat{\varphi}$	0.16347 (2)	0.24145 (7)	0.17842 (4)	$0.24070^{(0)}$	0.12275	0.27309	0.16427	0.19753 (5)
80		λ 2	0.00614 (7)	0.00590	0.00549	0.00646	0.00224	0.00582 (4)	0.00513 (2)	0.00600 (0)
		$\theta$	0.50150 (8)	0.33442 (3)	0.35764 {3}	0.39078 [0]	0.06108 (1)	0.33595 (4)	0.33240 (2)	0.39877 {7}
		θ	0.37274 [8]	0.28555 (2)	0.34004 {/}	0.29396 (3)	0.08158 (1)	0.31620 <sup>{4</sup> }	0.32734 (0)	0.31845 (5)
	MREs	$\hat{\varphi}$	0.43977 {3}	0.54161 (6)	0.45690 <sup>{4</sup> }	0.54328 {/}	0.36821 {1}	0.56628 (8)	0.43472 {2}	0.47912 (5)
		λ	0.12992 1/3	0.12509 (3)	0.11952 (5)	0.13116 [8]	0.07620 {1}	0.12061 (4)	0.11448 (2)	0.12607 [0]
		θ	0.48126	0.36089 (2)	0.38260 {3}	0.39633 (0)	0.15279 {1}	0.36296 (5)	0.36362 (4)	0.40027 {/}
		$\sum_{n} Ranks$	77 [8]	47 (5)	57 (4)	71 {/}	<u>12 (1)</u>	<b>59</b> (3)	39 {2}	70 (0)
		θ	1.18071 (8)	1.13329 (4)	1.15680 (0)	1.15101 (3)	0.12881 (1)	1.09310 {2}	1.11947 (3)	1.17399 {/}
	BIAS	$\hat{\varphi}$	0.21288 (2)	0.27282 (6)	0.22827 (4)	0.28207 {/}	0.16810 {1}	0.31936 (8)	0.21687 (3)	0.23691 (3)
		λ â	0.05455	0.05483	0.05139 (4)	0.05648 (8)	0.02373 (1)	0.04586 (2)	0.05000 (3)	0.05607 (7)
		θ	0.50995	0.46810 (5)	0.45161 (4)	0.48892 (7)	0.11088 (1)	0.36948 (2)	0.42847	0.48519 (0)
		θ	1.65210 (0)	1.59933	1.54359 (4)	1.65457 (7)	0.11230	1.41356 (2)	1.45126 (3)	1.65710 (8)
• • • •	MSEs	$\hat{\varphi}$	0.06995 (2)	0.11564	0.08167 (4)	0.12209 (7)	0.04607 (1)	0.16683	0.07397	0.08772 (3)
200		λ â	0.00441 {3}	0.00457 (6)	0.00403 (4)	0.00483 (8)	0.00088 {1}	0.00341 {2}	0.00386 (3)	0.00477 {7}
		θ	0.38578 [8]	0.35447 {3}	0.32071 (4)	0.38130 {/}	0.02076 {1}	0.24604 (2)	0.29346	0.38069 (0)
		θ	0.33734 (8)	0.32380 (4)	0.33052 (6)	0.32886 (3)	0.03680 {1}	0.31232 {2}	0.31985 (3)	0.33543 {/}
	MREs	$\hat{\varphi}$	0.28384 (2)	0.36376 [6]	0.30437 {4}	0.37610 {7}	0.22413 {1}	0.42582 (8)	0.28917 (3)	0.31587 {3}
		λ â	0.10910 (3)	0.10967 (0)	0.10278 (4)	0.11295	0.04/45	0.09173 (2)	0.10000 (3)	0.11214 (7)
		θ	0.40796	0.37448 (5)	0.36129	0.39114	0.08871	0.29559 (2)	0.34277	0.38815
		$\sum_{n} Ranks$	<b>67</b> 107	<b>64</b> (5)	52 (4)	83 [0]	12 <sup>(1)</sup>	<b>42</b> <sup>(3)</sup>	36 (2)	76 17
		ϑ ^	0.99609 (2)	1.12481 (6)	$1.05940^{(3)}$	1.12860 (7)	0.04256	1.04542 (*)	1.03887 (3)	1.140/9 (6)
	BIAS	$\varphi_{\hat{\lambda}}$	0.14496 (2)	0.18941	0.15844 (4)	0.19306 (7)	0.10411	0.23525 (8)	0.15349	0.16453 <sup>(3)</sup>
		л â	0.04309 (3)	0.05010 (7)	0.04425 (5)	0.04998 [0]	$0.014/4^{(1)}$	0.03603 (2)	0.04334 (4)	0.05109 (8)
		θ ô	0.38900 (4)	0.446/0 (6)	0.39875	0.45266 (7)	0.06374	0.31824 (2)	0.38560	0.45867
		ϑ ^	1.21294 (2)	1.50352 [6]	1.31440 <sup>(3)</sup>	1.52047 [7]	0.02893	1.25338 [3]	1.26686 (3)	1.51668 (7)
500	MSEs	Ŷ	0.03263 (2)	0.05612	0.03902	0.05818 (7)	0.01736	0.09221	0.03639	0.04120
500		л â	0.00278 (3)	0.00378 (0.5)	0.00295	0.00378 (0.3)	0.00034	0.00220 (2)	0.00283 (4)	0.00400 [8]
		θ ô	0.23678 143	0.31580	0.25000 (5)	$0.32430^{(7)}$	0.00703 (1)	0.18536 (4)	0.23280 (3)	0.33477 (8)
		ϑ ^	0.28460 [2]	0.32137 (0)	0.30269	0.32246 1/3	0.01216 (1)	0.29869 (*)	0.29682	0.32594 [8]
	MREs	$\varphi$	$0.19328^{12}$	0.25255	0.21126 (*)	0.25741 17	0.13881 (1)	0.31367	0.20466 13	0.21938
		л õ	0.08617	0.10020	0.08850	0.09997	0.02948	$0.07207^{12}$	0.08669	0.10217
		θ ΣΡΙ	0.31120	U.33/36 107	0.31900	0.30213 1/7	0.05100	U.23439 127	0.30848	0.30094
		> Kanks	33 121	<b>/4.3</b>	5/ 10	02.3 U	12 11	47 (*)	40 <sup>137</sup>	<b>80</b> (6)

n	Est.	Est. Par.	MLEs	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADEs
		ŵ	0 98923 {7}	0 64331 {3}	0 95672 {6}	0 63413 {2}	0 44818 {1}	1 04755 {8}	0 90857 {5}	0 78916 {4}
		ŵ	0.95367 <sup>{6}</sup>	0.90318 {4}	0.88553 {3}	0.95450 {7}	0.31253 {1}	0.97463 [8]	0.85825 <sup>{2}</sup>	0.93652 {5}
	BIAS	λ	0.50387 [8]	0.29370 {2}	0.38814 {5}	0.31558 {3}	0.23477 {1}	0.41787 {7}	0.38863 (6)	0.36144 43
		$\hat{ heta}$	0.83605 {8}	0.71444 {3}	0.81238 {7}	0.68991 {2}	0.62151 {1}	0.74404 {5}	0.75604 {6}	0.71561 {4}
		$\hat{artheta}$	1.35323 7	0.71389 {2}	1.18225 (6)	0.73718 {3}	0.37586 {1}	1.57224 {8}	1.13154 {5}	1.02701 {4}
	MOD	$\hat{arphi}$	1.18857 <sup>{8}</sup>	1.06788 {4}	1.02750 {3}	1.18599 {7}	0.87512 {1}	1.15778 {6}	0.96462 {2}	1.13400 {5}
20	MSEs	λ	0.36432 {8}	0.14714 {2}	0.22539 {5}	0.17712 {3}	$0.09172^{\{1\}}$	0.28638 {7}	0.23923 6	0.22267 4
		$\hat{ heta}$	0.77865 {8}	0.64926 {4}	0.77561 {7}	0.60519 {2}	0.52925 {1}	0.67762 {5}	0.68866 {6}	0.64046 {3}
		$\hat{artheta}$	0.28264 73	0.18380 {3}	0.27335 {6}	0.18118 {2}	0.12805 {1}	0.29930 [8]	0.25959 {5}	0.22547 4
	MDE	$\hat{arphi}$	0.47683 863	0.45159 43	0.44276 {3}	0.47725 {7}	0.40539 {1}	0.48731 {8}	0.42912 {2}	0.46826 {5}
	MRES	λ	0.33591 88	0.19580 {2}	0.25876 {5}	0.21038 {3}	0.15651 {1}	0.27858 {7}	0.25908 68	0.24096 44
		$\hat{ heta}$	0.33442 {8}	0.28578 {3}	0.32495 7	0.27596 {2}	$0.24860^{\{1\}}$	0.29762 {5}	0.30242 863	0.28624 4
		$\sum Ranks$	<b>89</b> <sup>{8}</sup>	<b>36</b> <sup>{2}</sup>	<b>63</b> <sup>{6}</sup>	<b>43</b> <sup>{3}</sup>	<b>12</b> {1}	<b>82</b> <sup>{7}</sup>	<b>57</b> <sup>{5}</sup>	<b>50</b> <sup>{4}</sup>
		$\frac{\overline{\hat{\vartheta}}}{\hat{\vartheta}}$	1.00512 {8}	0.76116 {3}	0.98814 {7}	0.76068 {2}	0.22490 {1}	0.94177 {5}	0.95845 {6}	0.83406 {4}
		$\hat{arphi}$	0.56247 {3}	0.64453 863	0.57353 {4}	0.64522 {7}	0.46925 {1}	0.76277 {8}	0.55915 {2}	0.63949 {5}
	BIAS	λ	0.45485 {8}	0.31115 {2}	0.36895 863	0.33000 {3}	0.11306 {1}	0.39112 {7}	0.36093 (5)	0.34405 44
		$\hat{ heta}$	0.76702 {8}	0.62063 {3}	0.68635 {7}	0.62598 {4}	0.33066 {1}	0.60384 {2}	0.66063 <sup>{6}</sup>	0.63130 {5}
		$\hat{artheta}$	1.24539 {7}	0.92842 {2}	1.16588 {6}	0.94580 {3}	0.13919 {1}	1.25646 {8}	1.12647 {5}	1.05555 {4}
	MOD	$\hat{arphi}$	0.45689 {3}	0.57098 863	0.46410 {4}	0.57244 {7}	0.32482 {1}	0.71835 {8}	0.44181 {2}	0.55971 {5}
80	MSEs	λ	0.28629 {8}	0.17260 {2}	0.20504 {5}	0.19582 {3}	0.02220 {1}	0.25122 {7}	0.20232 {4}	0.20735 68
		$\hat{ heta}$	0.68706 {8}	0.51170 {3}	0.58072 {7}	0.52071 {4}	0.16747 {1}	0.49400 {2}	0.55280 {6}	0.52733 {5}
		$\hat{artheta}$	0.28718 {8}	0.21747 {3}	0.28233 {7}	0.21734 {2}	0.06426 {1}	0.26908 {5}	0.27384 {6}	0.23830 {4}
		$\hat{arphi}$	0.28123 {3}	0.32226 [6]	$0.28677^{\{4\}}$	0.32261 {7}	0.23462 {1}	0.38138 {8}	0.27957 {2}	0.31974 {5}
	MREs	λ	0.30323 (8)	0.20744 {2}	0.24597 863	0.22000 {3}	0.07538 {1}	0.26075 {7}	0.24062 {5}	0.22937 43
		$\hat{ heta}$	0.30681 {8}	0.24825 {3}	0.27454 {7}	0.25039 43	0.13226 {1}	0.24153 223	0.26425 863	0.25252 {5}
		$\sum Ranks$	<b>80</b> <sup>{8}</sup>	<b>41</b> <sup>{2}</sup>	<b>70</b> <sup>{7}</sup>	<b>49</b> <sup>{3}</sup>	<b>12</b> <sup>{1}</sup>	<b>69</b> <sup>{6}</sup>	<b>55</b> <sup>{4}</sup>	<b>56</b> <sup>{5}</sup>
		$\hat{\vartheta}$	0.98460 [6]	0.86500 {3}	1.00674 {7}	0.86298 {2}	0.12154 {1}	1.05269 [8]	0.98185 {5}	0.89235 4
		$\hat{arphi}$	0.38859 {2}	0.45911 853	0.41129 43	0.46722 {6}	0.31253 {1}	0.67732 {8}	0.40329 {3}	0.46920 77
	BIA5	λ	0.39873 [8]	0.31023 23	0.35370 {6}	0.32546 {3}	$0.06881^{\{1\}}$	0.37938 7	0.34126 (5)	0.32900 44
		$\hat{ heta}$	0.68985 {8}	0.57483 {3}	0.64129 7	0.59868 {5}	0.21029 {1}	0.55195 23	0.61763 863	$0.59808^{4}$
		$\hat{artheta}$	1.11626 863	1.01791 {2}	1.13048 73	1.02064 {3}	$0.05498^{\{1\}}$	1.30905 [8]	1.09868 [5]	1.05550 {4}
	MCE	$\hat{arphi}$	0.23053 23	0.31234 {5}	0.25573 44	0.32276 {6}	$0.15678^{\{1\}}$	0.58268 {8}	0.24649 33	0.32421 7
200	MSES	λ	0.22103 7	0.16583 23	0.18036 {5}	0.17985 43	0.00837 {1}	0.22479 [8]	0.17270 {3}	0.18259 863
		$\hat{ heta}$	0.58585 {8}	0.45832 {3}	0.51809 73	0.48976 {5}	$0.07171^{\{1\}}}$	0.44322 {2}	0.49209 863	$0.48968^{\{4\}}$
		$\hat{artheta}$	0.28132 863	0.24714 {3}	0.28764 {7}	0.24657 23	0.03472 {1}	0.30077 [8]	0.28053 (5)	0.25496 43
	MDE	$\hat{arphi}$	0.19429 {2}	0.22956 {5}	0.20564 4	0.23361 [6]	0.15627 11	0.33866 [8]	0.20165 33	0.23460 {7}
	WIKES	λ	0.26582 {8}	0.20682 {2}	0.23580 163	0.21697 33	0.04587 113	0.25292 7	0.22751 [5]	0.21933 44
		$\hat{ heta}$	0.27594 88	0.22993 33	0.25652 {7}	0.23947 53	0.08412 {1}	0.22078 23	0.24705 68	0.23923 43
		$\sum Ranks$	<b>71</b> <sup>{6.5}</sup>	<b>38</b> <sup>{2}</sup>	<b>71</b> <sup>{6.5}</sup>	<b>50</b> <sup>{3}</sup>	<b>12</b> <sup>{1}</sup>	<b>76</b> <sup>{8}</sup>	<b>55</b> <sup>{4}</sup>	<b>59</b> <sup>{5}</sup>
		$\hat{artheta}$	0.93304 {5}	0.93263 44	0.97539 77	0.92961 {3}	0.05348 {1}	1.08240 [8]	0.95522 8	0.92527 <sup>{2}</sup>
	BIAS	$\hat{arphi}$	0.27572 {2}	0.32936 [5]	0.29210 43	0.33160 68	0.19113 {1}	0.58412 [8]	0.28766 {3}	0.33820 {7}
	DIT IS	Â	0.34882 7	0.31976 {2}	0.33838 [6]	0.33117 {5}	0.03908 {1}	0.35124 [8]	0.32736 {3}	0.32969 4
		$\hat{ heta}$	0.61892 <sup>{8}</sup>	0.57580 {3}	0.60923 7	0.59341 <sup>{6}</sup>	0.12272 {1}	0.51040 {2}	0.59246 <sup>{5}</sup>	0.58416 <sup>{4}</sup>
		$\hat{artheta}$	0.99386 {2}	1.04383 863	1.04849 7	1.04352 5	0.00732 {1}	1.28318 <sup>{8}</sup>	1.01861 <sup>{3}</sup>	1.03056 <sup>{4}</sup>
	MSEs	$\hat{arphi}$	0.11591 <sup>{2}</sup>	0.16651 <sup>{5}</sup>	0.13079 43	$0.16807^{\{6\}}$	0.06394 {1}	0.45248 {8}	0.12739 33	0.17250 7
500	WIGES	λ	$0.17000^{7}$	0.15789 {3}	0.15993 4	0.16663 5	0.00250 {1}	0.18698 [8]	0.15186 {2}	0.16915 863
		$\hat{ heta}$	$0.48714^{8}$	0.45146 {4}	0.47250 863	0.47497 <sup>{7}</sup>	0.02504 {1}	0.39915 <sup>{2}</sup>	0.45105 3	0.47078 {5}
		$\hat{\vartheta}$	0.26658 (5)	0.26647 4	0.27868 7	0.26560 {3}	0.01528 {1}	0.30926 [8]	0.27292 8	0.26436 {2}
	MRE	$\hat{arphi}$	0.13786 23	0.16468 [5]	0.14605 44	0.16580 863	0.09557 {1}	0.29206 [8]	0.14383 33	0.16910 <sup>{7}</sup>
		λ	0.23255 [7]	0.21318 [2]	0.22559 [6]	0.22078 [5]	0.02606 {1}	0.23416 [8]	0.21824 {3}	0.21980 [4]
		$\theta$	0.24757 88	0.23032 {3}	0.24369 7	0.23736 [6]	0.04909 {1}	0.20416 {2}	0.23698 (5)	0.23367 43
		$\sum Ranks$	<b>63</b> {5.5}	<b>46</b> <sup>{3}</sup>	<b>69</b> <sup>{7}</sup>	<b>63</b> {5.5}	<b>12</b> <sup>{1}</sup>	<b>78</b> <sup>{8}</sup>	<b>45</b> <sup>{2}</sup>	<b>56</b> <sup>{4}</sup>

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**Table 6.** Simulation results for  $\boldsymbol{\eta} = (\vartheta = 4.5, \varphi = 0.75, \lambda = 0.5, \theta = 1.25)^{\mathsf{T}}$ .

	Б. (	E / D		LOE	NH OF	CDUNE	MDGE	DOE		DADE
n	Est.	Est. Par.	MLES	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADES
		$\hat{artheta}$	1.25197 [8]	0.45438 {3}	1.06521 {7}	0.44888 {2}	0.20680 {1}	0.83805 {5}	1.02528 [6]	0.62513 {4}
		$\hat{arphi}$	0.58123 {5}	0.60203 {7}	0.57995 <sup>{4}</sup>	0.61278 {8}	0.16575 {1}	0.57749 {3}	0.53508 {2}	0.58165 {6}
	BIAS	λ	0.08854 {8}	0.07652 {2}	0.08227 {7}	0.07831 {3}	0.06474 {1}	0.08065 {6}	0.07939 {4}	0.07980 {5}
		$\hat{ heta}$	0.66759 {8}	0.42262 {2}	0.53041 73	0.44839 {3}	0.32466 {1}	0.51024 153	0.52797 863	0.48959 4
		$\hat{\vartheta}$	2.41625 (8)	0.60512 {2}	1.67131 {6}	0.67908 {3}	0.18285 {1}	1.58289 {5}	1.74298 {7}	1.01407 44
	MOL	$\hat{arphi}$	0.42148 {3}	0.46192 77	0.43565 {5}	0.46874 {8}	0.38218 {1}	0.43432 43	0.38271 {2}	0.43679 863
20	MSEs	λ	0.01072 {8}	0.00870 {2}	0.00963 77	0.00906 {3}	0.00620 {1}	0.00950 863	0.00923 (5)	0.00919 43
		$\hat{ heta}$	0.58894 {8}	0.28693 23	0.40885 863	0.32249 {3}	0.16767 {1}	0.38646 {5}	0.41599 {7}	0.36863 43
		$\hat{artheta}$	0.27821 {8}	0.10097 33	0.23671 {7}	0.09975 23	$0.04596^{\{1\}}$	0.18623 5	0.22784 8	0.13892 4
	MDE-	$\hat{arphi}$	0.77497 <sup>{5}</sup>	0.80271 {7}	0.77327 43	0.81704 {8}	0.70903 11	0.76998 33	0.71344 {2}	0.77554 463
	MKES	λ	0.17707 <sup>{8}</sup>	0.15303 23	0.16453 7	0.15662 {3}	0.12948 11	0.16130 863	0.15879 43	0.15960 {5}
		$\hat{ heta}$	0.53407 88	0.33809 23	0.42432 77	0.35871 {3}	0.25973 113	0.40819 853	0.42237 863	0.39167 43
		$\sum Ranks$	<b>85</b> <sup>{8}</sup>	<b>41</b> <sup>{2}</sup>	<b>74</b> <sup>{7}</sup>	<b>49</b> <sup>{3}</sup>	<b>12</b> <sup>{1}</sup>	<b>58</b> <sup>{6}</sup>	<b>57</b> <sup>{5}</sup>	<b>56</b> <sup>{4}</sup>
		$\hat{artheta}$	1.62748 [8]	0.91367 23	1.37823 7	0.98139 33	0.11224 {1}	1.13720 {5}	1.31803 68	1.11805 44
	BIAS	$\hat{arphi}$	0.31837 {3}	0.39968 77	0.34474 43	0.40380 [8]	0.27352 {1}	0.37466 [6]	0.31821 23	0.35079 {5}
	DIT IS	Â	0.07169 <sup>{8}</sup>	0.05804 {3}	0.06253 7	0.06113 68	0.03286 {1}	0.05699 <sup>{2}</sup>	0.05980 <sup>{5}</sup>	0.05870 <sup>{4}</sup>
		$\hat{ heta}$	0.64853 <sup>{8}</sup>	0.39956 {2}	0.49396 <sup>{7}</sup>	0.44278 43	0.16292 {1}	0.43570 {3}	0.47512 68	0.45276 (5)
		$\hat{artheta}$	3.14585 <sup>{8}</sup>	1.58782 {2}	2.33956 {7}	1.79431 {3}	$0.08679^{\{1\}}$	1.96177 <sup>{4}</sup>	2.22845 [6]	2.05141 {5}
	MSEs	$\hat{arphi}$	0.15254 {2}	0.23761 <sup>{7}</sup>	0.18156 {4}	0.24031 [8]	0.12335 {1}	0.22013 [6]	0.15798 {3}	0.19064 {5}
80	101010	λ	0.00707 [8]	0.00520 <sup>{4</sup> }	0.00579 {7}	0.00566 [6]	0.00171 {1}	0.00514 {2}	0.00541 {5}	0.00519 {3}
		$\hat{\theta}$	0.56658 [8]	0.28713 {2}	0.38674 {7}	0.34041 <sup>{4}</sup>	0.04197 {1}	0.32550 {3}	0.36722 [6]	0.35156 {5}
		θ	0.36166 {8}	0.20304 {2}	0.30627 {7}	0.21809 {3}	0.02494 {1}	0.25271 {5}	0.29289 [6]	0.24846 4
	MREs	$\hat{arphi}$	0.42449 {3}	0.53291 {/}	0.45966 [4]	0.53840 [8]	0.36469 {1}	0.49955 [6]	0.42429 {2}	0.46771 {5}
		λ	0.14337 {8}	0.11609 (3)	0.12506 {/}	0.12226 (6)	0.06572 {1}	0.11399 {2}	0.11961 (5)	0.11740 (4)
		$\theta$	0.51882	0.31965	0.39517	0.35422	0.13034	0.34856	0.38010	0.36221 (3)
		<u>}</u> Kanks	<b>80</b> (8)	<b>43</b> <sup>(-)</sup>	1 40301 {7}		0.04307 {1}	<b>47</b> <sup>(8)</sup>	<b>30</b> (8)	<b>54</b> (5)
		ô	$0.20186^{\{2\}}$	$0.26830^{7}$	0 21385 <sup>{4}</sup>	$0.26575^{\{6\}}$	0.16575 {1}	$0.27408^{\{8\}}$	0 20558 {3}	0.23166 <sup>{5</sup> }
	BIAS	φ ĵ	0.06056 <sup>{8}</sup>	0.05260 {3}	0.05573 {7}	0.20575	0.02038 {1}	0.27400 $0.04542^{\{2\}}$	0.05309 {4}	0.05470 {5}
		Â	0.54563 <sup>{8}</sup>	0.03200 $0.42847^{\{3\}}$	0.03373 0.47838 <sup>{7}</sup>	0.05571 0.45810 <sup>{5}</sup>	0.02050	$0.36816^{\{2\}}$	0.03307	0.0570
		พิ	2 63270 <sup>{8}</sup>	2 08923 <sup>{3}</sup>	2 37558 {7}	2 20468 <sup>{5}</sup>	$0.02535^{\{1\}}$	1 86391 <sup>{2}</sup>	2 18628 <sup>{4}</sup>	2 32684 <sup>{6}</sup>
		ô	$0.06343^{\{2\}}$	$0.11259^{\{7\}}$	$0.07126^{\{4\}}$	$0.11064^{\{6\}}$	0.02333	0 12626 <sup>{8}</sup>	0.06649 <sup>{3}</sup>	0.08397 <sup>{5}</sup>
200	MSEs	Â	0.00531 <sup>{8}</sup>	0.00437 {3}	0.00472 {6}	0.00486 {7}	0.00066 {1}	0.00347 <sup>{2}</sup>	0.00441 <sup>{4}</sup>	0.00470 {5}
		$\hat{\theta}$	0.44122 {8}	0.33416 {3}	0.37078 {5}	0.37139 [6]	0.01519 {1}	0.25838 {2}	0.33754 {4}	0.37704 <sup>{7}</sup>
		$\hat{\vartheta}$	0.32920 {8}	0.26063 {2}	0.31198 {7}	0.26954 4	0.00957 {1}	0.26211 {3}	0.29645 (6)	0.28705 (5)
	1055	$\hat{arphi}$	0.26915 {2}	0.35774 {7}	0.28513 {4}	0.35434 {6}	0.22100 {1}	0.36544 {8}	0.27411 {3}	0.30888 {5}
	MREs	λ	0.12113 (8)	0.10520 {3}	0.11146 {7}	0.11142 {6}	0.04075 {1}	0.09084 {2}	0.10618 43	0.10941 {5}
		$\hat{ heta}$	0.43650 {8}	0.34278 {3}	0.38270 {7}	0.36648 {5}	$0.07797^{\{1\}}$	0.29453 22	0.35973 43	0.36986 863
		$\sum Ranks$	<b>78</b> <sup>{8}</sup>	<b>46</b> <sup>{3}</sup>	<b>72</b> <sup>{7}</sup>	<b>66</b> <sup>{6}</sup>	<b>12</b> <sup>{1}</sup>	<b>44</b> <sup>{2}</sup>	<b>49</b> <sup>{4}</sup>	<b>65</b> <sup>{5}</sup>
		$\hat{\vartheta}$	1.29672 {5}	1.27235 {3}	1.31237 [6]	1.32513 7	0.01519 {1}	1.15120 {2}	1.28994 4	1.34189 88
	BIAS	$\hat{arphi}$	0.13744 {2}	0.17802 863	0.14815 43	0.18664 7	0.10227 {1}	0.19392 88	0.14640 {3}	0.15865 5
	DIA5	Â	0.04970 {5}	0.05107 863	0.04913 43	0.05400 [8]	0.01264 {1}	0.03665 {2}	0.04825 33	0.05335 7
		$\hat{ heta}$	0.44102 53	0.44509 863	0.43050 43	0.47518 [8]	0.05962 {1}	0.31796 {2}	0.42160 33	0.46945 7
		$\hat{artheta}$	2.07266 [4]	2.19148 863	2.07839 53	2.34226 [8]	0.00344 {1}	1.67247 <sup>{2}</sup>	2.02330 33	2.31957 7
	MSEs	$\hat{arphi}$	0.02955 2	0.05030 <sup>{6}</sup>	0.03417 43	0.05288 {7}	0.01679 {1}	0.06276 <sup>{8}</sup>	0.03295 33	0.03874 {5}
500	1010125	λ	0.00377 <sup>{4}</sup>	0.00418 {6}	0.00378 {5}	0.00452 {7}	0.00025 {1}	0.00242 {2}	0.00363 {3}	0.00454 {8}
		$\hat{\theta}$	0.30987 5	0.34423 6	0.30505 {4}	0.37795 <sup>{8}</sup>	0.00564 {1}	0.19875 {2}	0.29242 {3}	0.37470 <sup>{7}</sup>
		$\vartheta$	0.28816 {5}	0.28275 {3}	0.29164 [6]	0.29447 7	0.00337 {1}	0.25582 {2}	0.28665 4	0.29820 [8]
	MREs	$\hat{arphi}$	0.18326 {2}	0.23736 [6]	0.19754 4	0.24885 {7}	0.13636 {1}	0.25856 [8]	0.19520 {3}	0.21153 (5)
		λ	0.09940 {5}	0.10214 (6)	0.09825 [4]	0.10801 <sup>{8}</sup>	0.02528 {1}	0.07330 {2}	0.09651 {3}	0.10671 {/}
		θ	0.35281 {5}	0.35607 [6]	0.34440 {4}	0.38014 {8}	0.04770 {1}	0.25437 {2}	0.33728 {3}	0.37556 {/}
		$\Sigma Ranks$	<b>49</b> <sup>{4</sup> }	<b>66</b> <sup>{0}</sup>	<b>54</b> <sup>{5}</sup>	<b>90</b> <sup>{8}</sup>	12 {1}	<b>42</b> <sup>{3}</sup>	<b>38</b> <sup>{2}</sup>	<b>81</b> {/}

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$\eta^{ op}$	n	MLEs	LSEs	WLSEs	CRVMEs	MPSEs	PCEs	ADEs	RADEs
	20	2	7	1	5	4	8	3	6
$(\hat{\vartheta} = 1.75 \ \hat{\alpha} = 0.5 \ \hat{J} = 0.25 \ \hat{\theta} = 1.25)$	80	3	7	2	6	1	5	4	8
$(b = 1.75, \psi = 0.5, \lambda = 0.25, b = 1.25)$	200	1.5	7	4	5	1.5	6	3	8
	500	2	5.5	4	5.5	1	7	3	8
	20	1	3	2	7	4	8	5.5	5.5
$(\hat{\vartheta} = 1.75, \hat{\alpha} = 2, \hat{\beta} = 1.5, \hat{\theta} = 2.5)$	80	1	3	4	6	2	8	5	7
$(b = 1.75, \psi = 2, \lambda = 1.5, b = 2.5)$	200	2	3	4.5	6.5	1	8	4.5	6.5
	500	3	4	5	6	1	8	2	7
	20	7	4	3	6	1	8	2	5
$(\hat{\vartheta} = 3.5, \hat{\alpha} = 0.75, \hat{\beta} = 0.5, \hat{\theta} = 1.25)$	80	8	3	4	7	1	5	2	6
$(b = 3.3, \psi = 0.73, \lambda = 0.3, b = 1.23)$	200	6	5	4	8	1	3	2	7
	500	2	6	5	7	1	4	3	8
	20	8	2	6	3	1	7	5	4
$(\hat{\vartheta} = 35, \hat{\alpha} = 2, \hat{\beta} = 15, \hat{\theta} = 25)$	80	8	2	7	3	1	6	4	5
$(\vartheta = 3.5, \hat{\varphi} = 2, \lambda = 1.5, \theta = 2.5)$	200	6.5	2	6.5	3	1	8	4	5
	500	5.5	3	7	5.5	1	8	2	4
	20	8	2	7	3	1	6	5	4
$(\hat{\vartheta} = 4.5, \hat{\alpha} = 0.75, \hat{\vartheta} = 0.5, \hat{\vartheta} = 1.25)$	80	8	2	7	6	1	3	5	4
$(b = 4.3, \psi = 0.75, \lambda = 0.5, b = 1.23)$	200	8	3	7	6	1	2	4	5
	500	4	6	5	8	1	3	2	7
	20	8	2	5	3	1	7	6	4
$(\hat{\vartheta} = 4.5, \hat{\alpha} = 0.75, \hat{\vartheta} = 0.5, \hat{\theta} = 1.25)$	80	6	2	4	3	1	8	7	5
$(v = 4.3, \varphi = 0.75, \lambda = 0.5, \theta = 1.23)$	200	6	2	4	3	1	7	8	5
	500	6	2	5	3	1	8	7	4
	20	1	6	8	2	7	5	4	3
$(\hat{\vartheta} = 0.25, \hat{\alpha} = 3.5, \hat{\beta} = 3, \hat{\theta} = 0.25)$	80	1.5	5	8	1.5	6	4	3	7
$(b = 0.25, \varphi = 5.5, \lambda = 5, b = 0.25)$	200	1	7	4	6	2	3	5	8
	500	2	8	6	7	1	3	4.5	4.5
	20	2	8	6	5	3	1	4	7
$(\hat{\theta} = 0.75, \hat{\alpha} = 0.5, \hat{\beta} = 4.25, \hat{\alpha} = 0.25)$	80	1.5	7	5	6	1.5	3	4	8
$(\vartheta = 0.75, \varphi = 2.5, \lambda = 4.25, \theta = 0.25)$		2	7	5	6	1	3	4	8
	500	2	8	5	6	1	3	4	7
$\sum Ranks$		133.5	143.5	160	164	54	176	130.5	190.5
Overall Rank		3	4	5	6	1	7	2	8

 Table 7. Partial and overall ranks of the classical estimation methods for several parametric values.

For each sample and each parametric combination, the EWFr parameters  $\vartheta$ ,  $\varphi$ ,  $\lambda$  and  $\theta$  are estimated using the eight estimators called MLEs, LSEs, WLSEs, MPSEs, CRVMEs, ADEs, PCEs, and RADEs. Simulated results are listed in Tables 2–6 which also indicate the ranks of each of the proposed estimators in each row, where the superscripts show the indicators, and the  $\sum Ranks$  illustrates the partial sum of ranks in each column for a particular sample size.

Table 7 lists the partial and overall ranks for all parametric combinations. From Tables 2–6, one can conclude that all methods of estimation illustrate the consistency property, that is, the MREs and MSEs decrease as the sample size increases, for all parametric combinations. Table 7 shows that the MPS approach has an overall score of 54, hence it outperforms other estimation methods. Therefore, we conclude and confirm the superiority of MPSEs for estimating the EWFr parameters. However, the ADEs and MLEs are approximately have a similar performance, where their overall scores are,

respectively, 130.5 and 133.5.

#### 5. Modeling medicine and engineering data

In this section, the flexibility and importance of the EWFr distribution in modeling real-life data are illustrated using two datasets from the medicine and engineering fields. The first dataset contains 128 observations and it refers to remission times (in months) for bladder cancer patients [23]. The second dataset contains 63 observations and it represents strengths for single carbon fibers of 10 mm of gauge lengths [24]. The fits of the EWFr distribution is checked and compared with some important extensions of the Fr distribution called the exponentiated-Fr (EFr) [6], beta-Fr (BFr) [7], odd Lomax-Fr (OLxFr) [25], Kumaraswamy Marshall–Olkin Fr (KMOFr) [26], gamma extended-Fr (GExFr) [27] and transmuted exponentiated-Fr (TEFr) [28] and Fr distributions.

We checked the competing distributions using some goodness-of-fit analytical measures such as AIC (Akaike information criterion), CAIC (consistent Akaike information criterion), HQIC (HannanQuinn information criterion), BIC (Bayesian information criterion),  $W^*$  (Cramér–Von Mises),  $A^*$  (Anderson–Darling), and KS (Kolmogorov–Smirnov) statistics with its PV (p-value).

The maximum likelihood (ML) estimates of the parameters of the fitted distributions, their standard errors (SEs), and the analytical measures are reported in Tables 8 and 9 for the cancer and gauge lengths data, respectively. The numbers in Tables 8 and 9 indicate that the EWFr distribution gives a superior fit over other competing models, since it has the lowest values for all measures and the largest PV.

Some plots including the PDF, CDF, and SF along with the PP plots of the EWFr model are displayed in Figure 3 for both datasets. The PP plots of all studied distributions are displayed in Figure 4 for the two datasets, respectively.



Figure 3. The fitted EWFr PDF, CDF, SF, and P-P plots of the EWFr distribution.

Model	Par	Estimates	(SEs)	AIC	CAIC	BIC	HQIC	$W^*$	$A^*$	KS	(PV)
	â	56 02662	(28 42845)								
	ů ô	0.46607	(36.42643)								
EWFr	$\hat{\varphi}$	0.40097	(0.22382)	827.52150	827.84670	838.92960	832.15660	0.01991	0.13402	0.03757	0.99713
	л ô	0.71623	(0.01713)								
	0	0.01794	(0.01207)								
	â	2 18260	(1.04763)								
	ŵ	625 38000	$(622\ 71234)$								
OLxFr	Ψ â	0.12210	(0.08023)	827.70270	828.02790	839.11080	832.33780	0.02380	0.16520	0.04060	0.98433
	ĥ	1 38560	(0.00023)								
	υ	1.38500	(0.17273)								
	$\hat{\vartheta}$	43.87750	(128.52731)								
	ŵ	0.52028	(0.41388)								
KMOFr	$\hat{\delta}$	0.01314	(0.00808)	831.05330	831.54510	845.31340	836.84730	0.26278	0.34950	0.04457	0.96120
	â	3.27359	(2.90929)								
	ĥ	19.93782	(47.58906)								
			(								
	â	1426.40000	(1440.19245)								
EFr	ĥ	0.26240	(0.02850)	830.85880	831.05240	839.41490	834.33520	0.06720	0.46640	0.04910	0.91754
	$\hat{ heta}$	46.24800	(22.87321)								
	-		(								
	$\hat{artheta}$	0.60970	(0.32261)								
	ŵ	36.60200	(19.48341)								
BFr	â	739.38700	(629.24021)	833.09830	833.42350	844.50640	837.73350	0.06900	0.47650	0.05540	0.82683
	ĥ	0.32240	(0.06044)								
	-		(,								
	$\hat{\vartheta}$	226.67000	(267.96324)								
CE-E-	$\hat{\varphi}$	91.93900	(133.94324)	940 21620	940 54150	951 62450	944 95150	0 1 4 2 2 0	0.05070	0.06640	0 62502
GEXFF	â	22.27100	(119.19139)	840.21030	840.34130	831.02430	844.83130	0.14320	0.93970	0.00040	0.62393
	$\hat{b}$	0.07090	(0.04132)								
	$\hat{artheta}$	3.25820	(0.40714)								
TFF-	$\hat{arphi}$	0.75210	(0.04224)	802 00150	802 00750	207 70560	804 21010	0 74420	1 5 1 6 1 0	0 14090	0.01254
IEFF	â	39.38721	(40.24321)	892.00150	892.09730	897.70300	694.51910	0.74450	4.34040	0.14080	0.01234
	$\hat{b}$	0.25243	(0.76041)								
Fr	$\hat{ heta}$	0.75208	(0.04242)	802 00150	802 00750	807 70560	80/ 31010	0 74432	1 54642	0 14070	0.01250
	Â	2.43109	(0.21928)	092.00130	092.09130	071.10500	074.31910	0.74432	7.34042	0.14079	0.01230

 Table 8. The parameters estimates of the competing distributions and goodness-of-fit measures for cancer data.

The estimates of the EWFr parameters under several estimation approaches and the goodness-of-fit measures for both datasets are listed in Tables 10 and 11, respectively. Based on the values of PV in Tables 10 and 11, the Anderson-Darling approach is recommended to estimate the EWFr parameters for cancer data, while the least-squares approach is recommended for gauge lengths data.

Table 9. The parameters estimates of the competing distributions and goodness-of-fit
measures for gauge lengths data.

Model	Par	Estimates	(SEs)	AIC	CAIC	BIC	HQIC	$W^*$	$A^*$	KS	( <b>PV</b> )
	ŵ	0.76563	(0.3765631)								
	ŵ	0.04349	(0.3074545)								
EWFr	â	145.67068	(208.851634)	119.68850	120.37820	128.26110	123.06020	0.04181	0.23988	0.06847	0.92920
	$\hat{\theta}$	4.60846	(1.2033681)								
			( ,								
	$\hat{\vartheta}$	5.14301	(9.9790719)								
OI "Fr	$\hat{arphi}$	5.82568	(7.6204045)	110.04170	120 62140	129 51420	122 21240	0.04741	0 26520	0.07621	0 85661
OLXFF	â	4.17374	(2.671917)	119.94170	120.03140	128.31430	125.51540	0.04741	0.20330	0.07031	0.83004
	$\hat{b}$	2.82338	(0.9308868)								
	$\hat{artheta}$	31946.70000	(837.7074)								
	$\hat{arphi}$	3.40405	(1.493079)								
KMOFr	$\hat{\delta}$	3.35291	(0.9682264)	121.90480	122.95750	132.62050	126.11940	0.04616	0.26108	0.07450	0.87551
	â	0.85381	(0.1934628)								
	$\hat{b}$	13724.50000	(32837.29)								
	â	2.58672	(2.2847576)								
EFr	$\hat{b}$	0.50304	(0.4780356)	122.19730	122.88700	130.76990	125.56900	0.07082	0.40707	0.08894	0.70136
	$\hat{\theta}$	6.25000	(10.0300775)								
	~										
	$\vartheta$	2.38748	(5.65075)								
BFr	$\hat{arphi}$	1.00168	(5.322252)	120 58420	121 27380	129 15670	123 95580	0.06008	0 32203	0.07953	0 82045
211	â	22.38231	(162.860747)	120.30 120	121.27500	129.10070	120.00000	0.00000	0.02200	0.07755	0.02010
	ĥ	27.37096	(295.790324)								
	â	2 41716	(0.500(00))								
	v ^	2.41/16	(8.509608)								
GExFr	$\varphi$	1.20859	(2.106142)	120.57980	121.26950	129.15240	123.95150	0.06064	0.32329	0.07962	0.81934
	a î	24.77784	(82.649275)								
	b	15.60163	(85.339807)								
	â	1 04639	(1.0110058)								
	ô	2 50000	(1.0119038) (0.5073505)								
TEFr	Ψ â	2.30000	(0.5975595)	121.04420	121.73390	129.61670	124.41580	0.06547	0.34378	0.07500	0.87047
	î	0.00255	(2.0300897) (1.2750272)								
	D	0.09255	(1.2730272)								
	Â	5,43392	(0.50788)								
Fr	λ	230.48644	(110.91778)	121.80430	122.00430	126.09060	123.49010	0.11497	0.64202	0.10013	0.55274
Fr	$\hat{b}$ $\hat{\theta}$ $\hat{\lambda}$	5.43392 230.48644	(2.6300897) (1.2750272) (0.50788) (110.91778)	121.80430	122.00430	126.09060	123.49010	0.11497	0.64202	0.10013	0.55274

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Methods	$\hat{artheta}$	$\hat{arphi}$	λ	$\hat{ heta}$	$W^*$	$A^*$	KS	PV
MLEs	56.93661	0.46697	0.71825	0.01794	0.01991	0.13402	0.03757	0.99713
LSEs	56.93660	0.55502	0.71932	0.01871	0.01850	0.12986	0.03334	0.99918
MPSEs	56.93660	0.46698	0.71750	0.01717	0.01916	0.13069	0.03617	0.99708
WLSEs	56.94615	0.51962	0.71884	0.01825	0.01815	0.12600	0.03251	0.99947
CRVMEs	56.93659	0.55064	0.71962	0.01891	0.01795	0.12584	0.03352	0.99910
ADEs	56.93668	0.47772	0.71887	0.01815	0.01763	0.12016	0.03167	0.99968
PCEs	57.47389	0.81500	0.72219	0.02155	0.03320	0.23254	0.03830	0.99367
RADEs	56.93805	0.56319	0.71967	0.01899	0.01830	0.12878	0.03385	0.99895

**Table 10.** The estimates of the EWFr parameters under several methods of estimation and analytical measures for cancer data.

 Table 11. The estimates of the EWFr parameters under several methods of estimation and analytical measures for gauge lengths data.

	â	^	â	â	<b>TT</b> 7+	4 *	IZO	<b>DI</b> 7
Methods	$\vartheta$	arphi	λ	$\theta$	$W^*$	$A^*$	KS	PV
MLEs	0.76563	0.04349	145.67068	4.60846	0.04181	0.23988	0.06847	0.92920
LSEs	0.69421	0.08110	145.67010	4.56661	0.04850	0.26973	0.06486	0.94154
MPSEs	0.75534	0.13947	145.66920	4.65631	0.05531	0.29551	0.07780	0.81897
WLSEs	0.71596	0.01530	145.66740	4.57637	0.04747	0.26675	0.06895	0.90999
CRVMEs	0.71398	0.00100	145.67030	4.57259	0.04761	0.26713	0.06733	0.92336
ADEs	0.72963	0.00010	145.67980	4.57988	0.04696	0.26562	0.07047	0.89642
PCEs	0.75075	0.02596	145.67020	4.59649	0.04775	0.26737	0.07311	0.87060
RADEs	0.74062	0.00010	145.66970	4.57722	0.04668	0.26516	0.07058	0.89538



Figure 4. The PP plots of the EWFr distribution and other distributions.

# 6. Conclusions

In this paper, we introduced a more flexible four-parameter model called the extended Weibull– Fréchet (EWFr) distribution. Its basic mathematical properties are explored. The EWFr parameters are estimated using eight classical estimation methods. The simulation results showed that the maximum product of spacings approach outperforms other considered methods based on overall ranks. The importance and flexibility of the EWFr distribution over some competing extensions of the Fréchet distribution are addressed by analyzing two real-life datasets from the medicine and engineering fields. The analytical measures showed that our EWFr model returned an adequate fit in comparison with other distributions.

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# **Conflict of interest**

The authors declare no conflict of interest.

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